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Is Child Labor Harmful?
The Impact of Working as a Child on Adult Earnings

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Abstract

This paper explores the question: is working as young laborer harmful to an individual in terms of adult outcomes in income? This question is explored through the utilization of a unique set of instruments that control for the decision to work as a child *and* the decision of how much schooling to acquire. These instruments are combined with two large household survey data sets from Brazil that include retrospective information on the child labor and schooling of working-age adults: the 1988 and 1996 PNADs. Estimations of the reduced form earnings model are performed first by using OLS without controlling for the potential endogeneity of child labor and schooling, and then by using a GMM estimation of instrumental variables models that include the set of instruments for child labor and schooling. The findings of the empirical investigations show that child labor has large negative impact on adult earnings for both male and female children even when controlling for schooling. In addition, the negative impact of starting to work as a child reverses at around age 14.

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I. Introduction

Child labor is widespread in today's world, the International Labour Organization estimates that 246 million of the world's children, or 16 percent, are child laborers, most of whom live in developing countries.[1] Recently, there has been a renewed interest in child labor issues, and this renewed interest has led to a series of studies that aim to understand the causes and consequences of child labor in order to guide appropriate policy responses (see Basu (1999), for a useful survey of the theoretical and empirical literature). Among the policy options discussed are banning child labor and/or sanctioning countries that allow the practice. These types of policy responses have been widely debated among economists (see e.g., Emerson and Knabb, 2004; Basu, 2002; Dessy and Pallage, 2001; Baland and Robinson, 2000; Dessy, 2000; Basu and Van, 1998). Most of these studies emphasize the trade-off between child labor and human capital accumulation to justify policy interventions, arguing that there are large negative consequences from child labor. Child labor, however, is a catch-all term that encompasses a diversity of activities and working conditions, thus the belief that child labor is detrimental to human capital accumulation, may or may not be *generally* true, and, if so, at what age does this adverse effect cease to exist, and does the initial occupation matter, are open questions.[2] Studying and providing robust estimates of the

[1] Child labor was also common in developed countries until fairly recently, see, e.g., Kruse and Mahony, 2000.

[2] It should be noted that there are some forms of child labor the are unequivocally bad: those that are detrimental to a child's health and well-being, those that involve indentured servitude or deny children their basic human rights, and those that involve psychological distress, to identify a few. Some of these

effects of starting to work as a child on adult earnings will allow future studies of child labor, and discussions of appropriate policy responses, to be informed by these efforts.

Despite the fact that there is a large and growing literature on child labor, this one fundamental question remains unanswered: does child labor harm participants? Though it has been assumed to be detrimental, the potential effects of child labor on adult earnings are twofold. On one hand, child labor can be detrimental through the hindering of the acquisition of formal education, both quantitatively and qualitatively, and causing irreparable damage to health, reputation or other things that effect adult human capital, which could lead to lower wages in the adult labor market. On the other hand, there are many reasons why one might expect that there can be positive pecuniary benefits to young labor: vocational training, learning by doing, general workplace experience as well as the potential for making contacts, learning job market strategies, etc. In other words, there are many reasons to expect that a young laborer can gain some human capital from their workplace experience (e.g., Horn, 1994). Thus the net effect of starting to work as a child is an empirical issue. Though virtually all studies of child labor assume it is harmful, there is as yet no reliable measure of the effects of working as a child on adult outcomes.

These effects will also likely depend on the particular type of labor the child undertakes because some jobs may lead to the acquisition of job specific human capital while others may not. For example it could be that a child that works as manual laborer in agriculture does not learn many skills that the adult labor market values. However, an child that works as a manual laborer in a blacksmith shop may learn many skills of the

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activities may not be detrimental to the adult earnings of the individual, but are indefensible nonetheless and we do not wish to suggest otherwise.

blacksmith trade that are valuable on the adult labor market. For instance, French (2002) finds that child workers in shoe manufacturing industries in Brazil have positive attitudes toward their jobs if their work is associated with more autonomous and self-directed tasks. Furthermore, child labor could be a way to finance education that an individual would not otherwise have access to, which, in turn, could lead to better outcomes for older child or adolescent workers (see, e.g., Akabayashi and Psacharopoulos, 1999; Psacharopoulos, 1997).

Thus there is a startling gap between what is known about the causes of child labor and what is known about the consequences of child labor.[3] This paper seeks to fill this gap by providing a detailed analysis of two large survey datasets that contain retrospective information on the child labor of subjects and current information on their incomes.

This study analyzes the effects of starting to work as a child on the adult earnings of working age Brazilian men and women. Specifically, there are two main questions we will seek to answer: One, what is the effect on adult earnings of starting to work as a child both including the effect on educational attainment and the effect over and above the impact on schooling? Two, are there differences in these effects depending on which child labor occupations, or types of work, a child enters when first starting to work?

There are two main reasons for the lack of prior studies of this type: the lack of good data and the confounding effects of potentially endogenous variables. The present study is able to overcome both limitations.

[3] Some important exceptions are Emerson and Portela Souza (2003) who find that adult earnings are negatively related to the age the individual entered the labor force in Brazil. Ilahim, Orazem and Sedlacek (2000) find that child labor reduces the educational attainment and expected adult earnings in Brazil, however their study is constrained by lack of good instruments. Beegle, et. al. (2004) find that child labor

The first limitation is particular to developing countries where data of a high enough quality and a proper set of variables is hard to find. The Brazilian government, however, had been a pioneer in the collection and dissemination of data and thus Brazil presents a source of data that is adequate to address this complex relationship. The data utilized in this paper come from two rounds of a very large household survey from Brazil that includes information on current working-age adults' attributes and incomes as well as retrospective information on the age at which they first started to work and the number of years they attended school. These two data sets provide a very large and rich dataset from which to perform the analysis.

The second limitation is common to all studies that try and estimate the human capital - earnings relationship, which can be traced back to the seminal research of Becker and Chiswick (1966), Chiswick (1974) and Mincer (1974). Because of the strong likelihood that there are unobserved attributes (e.g., ability) that effect both the schooling choice of an individual and the individual's adult earnings, estimates that do not attempt to address this issue are considered unreliable. Recent research into this relationship using US data has relied on the use of instrumental variables to overcome the confounding effects of these unobserved attributes (Carneiro and Heckman, 2002; Card, 2001; Card, 1995a; Card 1995b). The main drawback of this type of approach is that it demands a robust set of instruments for the schooling choice of an individual. What makes this approach particularly challenging in the context of child labor is that schooling and child labor are likely jointly determined, so a set of suitable instruments must include instruments for *both* choices. To assemble a rich enough set of

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hampers educational attainment observed 5 years later. Others also argue that adolescent workers are more likely to end up in dead-end jobs that hamper their human capital development (Spindel, 1985).

instruments, data on the number of primary, secondary and college-level schools in each Brazilian state for each year from 1933 to 1976 was collected, along with the number of teachers in each state and year. These are hypothesized to be correlated with the work and schooling decisions of children (regardless of who made these choices), yet uncorrelated with adult incomes (when netting out state and other confounding effects). Statistical tests confirm the validity of these instruments.

With these two obstacles overcome, it is possible to provide a very detailed and robust estimation of the effects of child labor on adult incomes both including the effect of lost education and the effect over-and-above the effect on education. These results should be of vital interest to researchers of child labor in their quest to understand both the causes and consequences of child labor. This study then proceeds to test for differences in these observed effects of starting to work as a child by the occupation choice the person made when they first started to work.

The rest of this paper is organized as follows. Section 2 describes the data utilized in this study. Section 3 discusses the empirical strategy employed to explore the central questions of the paper. Section 4 presents the results of the empirical estimations. Section 5 discusses the results and implications.

II. The Data

The main sources of data utilized in this study are two rounds of the *Pesquisa Nacional por Amostragem a Domicílio* (PNAD), from *Instituto Brasileiro de Geografia e Estatística* (IBGE), the Brazilian census bureau: 1988 and 1996. The PNAD is a yearly and nationally representative household survey (excepting the rural Amazon region)

similar to the Current Population Survey in the U.S. It covers close to one hundred thousand households and includes information on the demographic and labor market characteristics of the households. Additionally, and of particular utility for the present study, the 1988 and 1996 surveys obtain retrospective information from the household head and the spouse about the age they entered the labor market, their first occupation, the educational attainment of their parents as well as the occupation of their fathers when they (head and spouse) first entered the labor market.

Additional data on the number of primary schools, secondary schools and colleges by state and year come from the Brazilian Census.

Our sample consists of all adults who are between 25 and 55 years of age at each survey year. We exclude younger and older adults in an attempt to avoid potential selectivity bias of labor participation decisions. Most of these prime age workers are likely to work in the labor market. While this restriction of the sample is adequate for males in our sample (of whom over 95% work), it is most likely inadequate for females in our sample (of whom less than 50% work). For females, we perform the analysis both with and without controls for participation decisions. For the estimation with controls we use the Heckman procedure to address the selection issues.[4] We perform all analyses on three samples: the 1988 and 1996 PNADs and a combined sample of both in which we control for cohort effects. The combined sample consists of around 50,000 complete observations for women and around 72,000 complete observations for men. The basic statistics of the combined sample are presented in Table A1 of the Appendix.

[4] Number of children is the additional variable used to control for the female labor market participation decision.

Before discussing the regression analysis, it is informative to show the distribution of the age started to work, the schooling attainment, and the log-earnings of the combined sample. The first column of Tables 1a and 1b show the distribution of the age started to work for males and females, respectively. We divide the age started to work into four groups: those that reported age started to work at nine years old or below, between ten and thirteen years old, between fourteen and seventeen years old, and those at eighteen years old or above. Around twenty percent of prime age male workers reported that they started to work at nine years old or below for both surveys and around forty percent of them started to work between ten and thirteen years old. For females, thirteen to fourteen percent of prime age women started to work at nine years old or below and around thirty percent of them started at age between ten and thirteen. Also, there a greater proportion of females did not start to work until adulthood compared to males. Around twelve to fifteen percent of male workers started to work at age eighteen or above, whereas around thirty percent of females started to work in this age range. Though these figures are available only for those who currently participate in the labor market, it seems that, in general, females start to work at an older age than do males.

Tables 1a and 1b also show the distribution of schooling attainment and log-earnings by each age-started-to-work group. We divide individuals into five educational attainment groups: illiterates, some primary or completed primary education, some secondary or completed secondary education, some high school or completed high school education, and some college or completed college education. Both tables reveal that individuals who started to work earlier in life have lower educational attainments and lower earnings as adults. For educational attainment, for instance, 26 percent of all

prime-age males that started to work at age nine or below are illiterate and two percent of them have some college or completed college education in 1988. Conversely, five percent of all prime-age males that started to work at eighteen years old or above are illiterates and 33 percent of them have some college or completed education in 1988. Similar patterns hold for females as well. Thus, Tables 1a and 1b appear to show that there is a direct relationship between the age and individual starts working and their educational attainment and adult earnings. Assuming a causal relationship between age-started-to-work and adult earnings either indirectly (through education) or directly (through experience) we can test if this positive relationship holds when controlling for other observable characteristics.

III. Empirical Strategy

In the typical Micerian framework of the effect of schooling on adult earnings in the high income country context, the discussion of the empirical issues usually begins with a presentation of a standard two equation system that describes schooling (S_i) and log wages ($\ln Y_i$), for an individual i:

$$(1) \qquad S_i = X_i\delta + \nu_i ,$$

$$(2) \qquad \ln Y_i = X_i\gamma + S_i\beta + \upsilon_i .$$

In this case X_i is a vector of observed attributes of the individual and ν_i and υ_i are the random error terms that are assumed to be uncorrelated with X_i. The coefficient β is a measure of the 'returns to education,' or average returns to education if this varies across individuals if ν_i and υ_i are uncorrelated.

It is quite likely that schooling is correlated with the unobserved component of the log earnings equation, however, due to ability bias (see, e.g. Griliches (1977)), measurement error in schooling, or a systematic variation in the returns to schooling based on years of schooling (higher marginal returns in earlier years of schooling, see Card (1995), for example). Ability bias arises when individuals of high ability both acquire higher levels of schooling (because the returns are higher and/or the costs are lower) and earn higher wages in the adult labor market. If this is true for our sample, an estimation of the β coefficient will be biased upwards. Measurement error in schooling can also bias the results if it induces a negative correlation between the errors of the observed schooling and earnings, which would bias the estimate of β downward. Finally, if individuals with lower levels of schooling have systematically higher returns to schooling (due to diminishing marginal returns to schooling in general) then estimates of β will also be downward biased.

The context of a low income country in which child labor is widespread presents another confounding effect: child labor itself. The decision to work as a child is likely correlated with the schooling decision and is also likely correlated with adult earnings. Fortunately, one aspect of child labor is observed: the age at which individuals first started to work. Therefore, in the low income country context, where child labor is widespread the schooling and child labor decision are both likely to effect adult returns to education and are likely correlated, a description of this process would involve a three equation system for and individual i:

(1) $S_i = X_i\delta + \nu_i$,

(2) $CL_i = X_i\alpha + \psi_i$,

(3) $$\ln Y_i = X_i\gamma + S_i\beta + CL_i\phi + \upsilon_i .$$

Where CL_i is the age at which the individual first started to work, and ψ_i is the unobserved random error term. Now, in order for ϕ to be a measure of the effect of starting to work at a certain age (or average if it varies across individuals), ψ_i and υ_i must be uncorrelated.

These error terms are likely correlated because of the same ability bias that causes high ability individuals to choose more schooling may cause those individuals to choose to start to work at an older age (biasing the coefficient estimate upward) or they may choose to start working at a younger age because ability may pay off in the child labor market as well as the adult labor market (biasing the coefficient estimate downward). Measurement error and a systematic variation in the impact of starting to work younger and the age at which the individual started to work are also sources of potential bias.

In this case, consistent estimates for the return to education and the effect of starting to work as a child can be obtained if there is a set of regressors Z_i that can be added to the vector X_i that affect schooling but do not affect earnings and that affect the age an individual starts to work but not earnings. This set of regressors must be sufficiently correlated with both schooling and the age started to work (i.e. have enough separate correlation with both variables that is separate from the correlation among the two variables), and sufficiently uncorrelated with adult earnings that they can be legitimately excluded from the earnings equation.

One set of variables that may fulfill this requirement are the number of primary schools, secondary schools and colleges per capita in the individual's state in the year that they are of the appropriate age to attend these schools. The presence of more schools

in the same state as the individual lowers the cost of attending school as travel costs are reduced and students are more likely to be able to live at home and attend school. Lower cost of education should increase investments in education, and cause delay in starting to work. However, if these variables are proxying for other things that affect adult earnings, like school quality for instance or the returns to education in a locally segmented labor market, then these variables will not serve as adequate instruments for the schooling choice and age at which the individual started to work. As it is possible that school quality is being proxied by the number of schools, we also employ a measure of the number of teachers per capita by state and year.

To test the model presented above, we estimate a series of OLS regressions and a series of GMM IV regressions in order to capture the effect of being an adolescent laborer on adult earnings. The first set of regressions will estimate the direct impact of being a young laborer on adult earnings. The second set of regressions will identify the first job occupations that are associated with higher or lower earnings conditional on having been an adolescent laborer. The third set of regressions will add the effect of having the same first job occupation and the father's occupation.

IV Estimation and Results

4.1 The Effect of Starting to Work as a Child

In order to estimate the effect of having been a child worker on current adult earnings, we start by estimating two separate earnings equations that include the age the individual first started to work variable and its square, the age of the individual and its square, indicator variables that equal one if the individual is classified as black and

another if the individual is classified as 'pardo,' or mixed race. Included in all estimations are measures of the father's and mother's education levels. For both, these are indicators for each level of education completed: lower primary, upper primary, secondary and college. For these estimations, an indicator variables that equals one if the individual resides in a rural area is included as well as indicators for the regions of brazil the individual currently resides in. The difference in the two separate earnings equations are that in the first estimations the years of schooling of the individual are not included and in the second set, the years of schooling are included. In all of the following cases we run these regressions for males and females separately.

We start by estimating the earnings model for the two survey years separately and estimate each first by OLS and then using the set of instruments described above in a GMM IV framework. The first set of regressions does not control for the individual's educational attainment. The fact that an individual worked during childhood or adolescence will likely mean that individual will have attained less education than a similar individual that did not work. So, as a first step, the coefficients of the young labor indicator variables when not controlling for education capture the expected forgone adult earnings of a young worker.

Tables 2 and 3 present the results for males and females in 1988, respectively. The first and fifth columns of each table show the coefficient estimates of the OLS and the second-stage of the IV regressions, respectively, when the individual's schooling variables *are not* included.[5] First, as we are interested in the young laborer status of the individual and its impact on his or her adult earnings, the coefficient estimates show that the older the individual enters the labor market, the higher are his or her earnings

(including the effect of the loss of education). For the IV estimation, the squared term is negative and significant, suggesting that this negative effect ceases at around age 14. Thus, there is a negative and significant impact on adult earnings if an individual started to work as a child at or below the age of 14, but that effect becomes positive for individuals who started to work at age 15 or above.

And the third and seventh columns present the result the coefficient estimates of the OLS and second-stage IV regressions, respectively, when the individual's schooling variables *are* included. Thus the coefficients estimates of the young laborer indicator variables reflect the effect on adult income of having been a young laborer over and above the loss of education. Here the age started to work coefficient estimate is still positive for both and its square is positive again for the IV estimation.

For all four estimations, the other coefficient estimates have the expected signs. Older individuals have higher earnings but this increases at a decreasing rate, black and pardo individuals have systematically lower earnings than white individuals, individuals in rural areas have lower earnings, and, the more educated the parents are, the greater the earnings of the individuals. These results hold for both male and female workers.

Tables 4 and 5 present the same estimations for the 1996 survey year with qualitatively the same results.

Tables 6 and 7 present the estimates for these same models for a pooled sample that included indicator for the 1988 survey year and indicators for cohorts: one for individuals born in the years 1933 to 1945, and another for individuals born in 1946-1958. The pooled sample estimations follow the same pattern of results as the preceding estimations. The '46-'58 cohort indicator variable estimate is positive and significant in

[5] All first-stage regression results are presented in the appendix.

all cases and, interestingly, individuals form the earlier sample, 1988, have systematically lower earnings, perhaps reflecting the growth of the Brazilian economy over the intervening years.

Together, these results suggest that there is indeed a negative impact of being a child labor both including the effect on educational attainment and over and above the impact on education. However, this effect seems to subside and turn positive at around age 15. From these exercises the picture that emerges is that though there is an important impact of adolescent labor on adult earnings through the trade-off with schooling, there is a strong impact over and above the effect on educational attainment.

4.2 The Role of Different Child Labor Activities

Since some activities that children may engage in when they work may have good vocational or other job training aspects to them, we next attempt to identify any particular activities that appear to have positive human capital. The distribution of the first job occupations for the pooled sample are given in Table 8. We construct five occupational categories from the three-digit occupation categories available in PNAD.[6] These categories are somewhat arbitrary at the margins since there do not exist very clear boundaries between the many occupations, but for the most part, capture the occupations generally associated with these activities. As Table 8 shows, the bulk of male child labor was devoted to agricultural activities. The majority of female child labor is concentrated in agricultural activities and in domestic work (part of services). Note that there are some changes in the child labor occupation distribution across the survey years.

[6] We are unable to estimate with enough precision, point estimates of more finely parsed occupation categories.

For instance, among male children that started to work at nine years old or below, more than 78 percent of them reported being involved in agricultural tasks in the 1982 survey, whereas only 55 percent of them reported to have started to work in agricultural activities in the 1996 survey. During the same time span, child laborers in civil construction, handcrafted activities, manual laborers and helpers in manufacturing increased in relative terms. This trend may reflect the increase in the urbanization in the last two decades in Brazil.

In order to estimate the impact of these specific child labor occupations on adult earnings, we estimate a series of IV models, similar to those presented previously that included schooling, but for each first job occupation separately. Also included are an indicator variable that equals one if the first job occupation of the individual was the same as the fathers occupation, and an interaction of this variable with age started to work. The coefficient estimates from these estimations are presented in Table 9. The key results here are that the effect of age started to work are similar to those not separated by occupation for commerce and transport and for services and others: positive and significant coefficient estimates for age started to work and negative and significant estimates for its square. Significant estimates are not obtained for the other three categories. This is not surprising for manufacturing and civil construction as they represent less than 10 percent and 5 percent of the sample, respectively. However the fact that that sign of the point estimate for the age started to work variable for the agricultural regression is negative, and the fact that agriculture accounts for almost 40% of the sample is intriguing. This suggests that there may be no adverse effect from starting to work as a child, over and above the impact on schooling, for those that

undertake agricultural activities. This is a particularly important result when one considers the fact that worldwide, 70 percent of child workers are estimated by the ILO to work in agriculture and related activities.

Thus there appears to be evidence that in some occupations entering as a child worker may not have adverse effects over and above the loss of education.

V. Conclusion

In this paper, we investigated the effect of starting to work as a child laborer on an individual's adult earnings. We find that child labor is associated with lower adult earnings, partly due to the trade-off associated with educational attainment and partly due to the effect over and above the impact on educational attainment, but that this negative effect appears to reverse around the age of 14.

Second, although there appears to be some decrease in adult earnings in general from child work beyond schooling, we find that for agricultural activities there appears to be no adverse effect. Particularly important for females is domestic work, which does not seem to harm the adolescent worker. Finally, we find that there are no gains for male workers associated with starting to work in the same occupations as their fathers.

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Table 1a: Schooling Distribution and Mean Log-Earnings by Age Started to Work
25 to 55 Year-Old Males

Age Started to Work	% Sample	Illiterate	Primary	Secondary	High School	College	Log-Earnings
				1988			
9 and Below	19.43	0.26	0.49	0.16	0.07	0.03	5.27
10 to 13	39.37	0.24	0.43	0.19	0.10	0.04	5.34
14 to 17	27.78	0.10	0.33	0.26	0.19	0.12	5.78
18 and Above	13.41	0.05	0.19	0.19	0.24	0.33	6.25
				1996			
9 and Below	18.68	0.23	0.42	0.23	0.09	0.03	5.64
10 to 13	36.44	0.19	0.37	0.26	0.13	0.05	5.75
14 to 17	30.04	0.07	0.25	0.32	0.25	0.12	6.12
18 and Above	14.83	0.04	0.13	0.23	0.30	0.30	6.37

Table 1b: Schooling Distribution and Mean Log-Earnings by Age Started to Work
25 to 55 Year-Old Females

Age Started to Work	% Sample	Illiterate	Primary	Secondary	High School	College	Log-Earnings
				1988			
9 and Below	12.55	0.293	0.486	0.162	0.044	0.015	4.291
10 to 13	29.84	0.263	0.455	0.173	0.080	0.028	4.364
14 to 17	25.69	0.099	0.291	0.237	0.231	0.142	5.010
18 and Above	31.93	0.057	0.205	0.181	0.277	0.280	5.322
				1996			
9 and Below	14.23	0.218	0.450	0.221	0.082	0.029	5.043
10 to 13	28.77	0.184	0.407	0.247	0.121	0.041	5.144
14 to 17	27.43	0.075	0.228	0.284	0.263	0.150	5.588
18 and Above	29.57	0.036	0.132	0.202	0.327	0.302	5.842

Table 2:

OLS and IV Estimates of Log-Earnings: 25-55 Year-Old Males 1988

Dependent Variables	OLS		OLS		IV 3		IV 4	
	Coeff.	Std. Error	Coeff.	Std. Error	Coeff.	Std. Error	Coeff.	Std. Error
Age Started to Work	0.033 ***	0.009	0.021 ***	0.006	1.495 ***	0.334	1.460 ***	0.322
Age Started to Work Squared	0.000	0.000	0.000 *	0.000	-0.045 ***	0.012	-0.050 ***	0.011
Years of Scholing			0.112 ***	0.002			0.205 **	0.082
Age	0.115 ***	0.010	0.102 ***	0.008	0.152 ***	0.020	0.131 ***	0.020
Age Squared	-0.001 ***	0.000	-0.001 ***	0.000	-0.002 ***	0.000	-0.001 ***	0.000
Back	-0.390 ***	0.022	-0.246 ***	0.020	-0.489 ***	0.046	-0.236 **	0.118
Pardo	-0.290 ***	0.013	-0.189 ***	0.012	-0.346 ***	0.031	-0.185 **	0.075
Father's Education								
Lower Primary	0.233 ***	0.012	0.062 ***	0.011	0.053	0.062	-0.188	0.098
Upper Primary	0.535 ***	0.033	0.155 ***	0.030	0.129	0.169	-0.333	0.216
Secondary	0.673 ***	0.033	0.192 ***	0.031	0.304	0.251	-0.228	0.294
College	0.767 ***	0.042	0.255 ***	0.041	0.507	0.381	0.069	0.375
Mother's Education								
Lower Primary	0.309 ***	0.012	0.116 ***	0.011	0.078	0.077	-0.160	0.107
Upper Primary	0.595 ***	0.034	0.219 ***	0.031	0.341 *	0.200	-0.075	0.238
Secondary	0.744 ***	0.034	0.320 ***	0.032	0.466 **	0.236	-0.010	0.273
College	0.796 ***	0.057	0.381 ***	0.055	0.603 **	0.290	0.220	0.307
Rural	-0.629 ***	0.014	-0.399 ***	0.014	-0.316 ***	0.123	-0.065	0.129
North	-0.225 ***	0.028	-0.162 ***	0.025	-0.312 ***	0.086	-0.124	0.111
Northeast	-0.300 ***	0.022	-0.213 ***	0.018	-0.301 ***	0.043	-0.121	0.085
South	-0.075 ***	0.026	-0.049 **	0.022	-0.045	0.035	-0.040	0.029
Center-West	-0.019	0.025	-0.032	0.020	0.181 ***	0.060	0.082	0.064
Constant	2.896 ***	0.193	2.744 ***	0.162	-8.434 ***	2.507	-7.917 ***	2.394
# Obs.	32,650		32,641		32,142		32,133	
R-Squared	0.366		0.481					
Hansen's J-Statistics					20.960		19.982	
Chi-Squared (P-value)					0.074		0.067	

Note: (1) *** statistically significant at 1%;** statistically significant at 5%;* statistically significant at 10%.

(2) The instruments are the nubmer of school by state and year.

Table 4:

IV Estimates of Log-Earnings: 25-55 Year-Old Males 1996

Dependent Variables	OLS		OLS		IV 3		IV 4	
	Coeff.	Std. Error	Coeff.	Std. Error	Coeff.	Std. Error	Coeff.	Std. Error
Age Started to Work	0.034 ***	0.007	0.021 ***	0.006	1.349 **	0.574	1.327 **	0.538
Age Started to Work Squared	0.000	0.000	-0.001 ***	0.000	-0.041 *	0.023	-0.042 **	0.021
Years of Scholing			0.105 ***	0.001			0.096	0.084
Age	0.105 ***	0.010	0.076 ***	0.008	0.123 ***	0.019	0.099 ***	0.030
Age Squared	-0.001 ***	0.000	-0.001 ***	0.000	-0.001 ***	0.000	-0.001 **	0.000
Back	-0.383 ***	0.020	-0.257 ***	0.018	-0.469 ***	0.050	-0.352 ***	0.118
Pardo	-0.311 ***	0.012	-0.211 ***	0.010	-0.324 ***	0.046	-0.239 ***	0.091
Father's Education								
Lower Primary	0.250 ***	0.012	0.078 ***	0.011	0.059	0.064	-0.068	0.134
Upper Primary	0.409 ***	0.023	0.096 ***	0.020	0.020	0.238	-0.187	0.304
Secondary	0.593 ***	0.029	0.199 ***	0.028	0.154	0.316	-0.101	0.392
College	0.858 ***	0.035	0.404 ***	0.031	0.595	0.623	0.332	0.647
Mother's Education								
Lower Primary	0.240 ***	0.011	0.081 ***	0.010	0.062	0.062	-0.060	0.125
Upper Primary	0.472 ***	0.024	0.169 ***	0.023	0.220	0.174	-0.006	0.270
Secondary	0.640 ***	0.030	0.259 ***	0.027	0.396	0.312	0.130	0.397
College	0.648 ***	0.044	0.229 ***	0.041	0.393	0.302	0.095	0.410
Rural	-0.619 ***	0.014	-0.407 ***	0.013	-0.206 *	0.119	-0.068	0.169
North	-0.209 ***	0.033	-0.141 ***	0.029	-0.220 ***	0.077	-0.146	0.092
Northeast	-0.298 ***	0.021	-0.212 ***	0.017	-0.273 ***	0.082	-0.182 *	0.101
South	-0.057 **	0.025	-0.025	0.020	0.046	0.058	0.059	0.055
Center-West	-0.030	0.025	-0.035	0.021	0.169 ***	0.049	0.145 ***	0.049
Constant	3.381 ***	0.191	3.595 ***	0.159	-6.610 *	3.666	-6.198 *	3.490
# Obs.	31,725		31,646		31,495		31,416	
R-Squared	0.376		0.489					
Hansen's J-Statistics					12.153		12.243	
Chi-Squared (P-value)					0.515		0.426	

Note: (1) *** statistically significant at 1%;** statistically significant at 5%;* statistically significant at 10%.

(2) The instruments are the nubmer of school by state and year.

Table 6:

OLS and IV Estimates of Log-Earnings: 25-55 Year-Old Males Pooled 1988 and 1996

Dependent Variables	OLS		OLS		IV 3		IV 4	
	Coeff.	Std. Error	Coeff.	Std. Error	Coeff.	Std. Error	Coeff.	Std. Error
Age Started to Work	0.032 ***	0.006	0.020 ***	0.004	1.937 ***	0.415	2.248 ***	0.411
Age Started to Work Squared	0.000	0.000	0.000 ***	0.000	-0.066 ***	0.017	-0.083 ***	0.018
Years of Scholing			0.109 ***	0.001			0.165 **	0.079
Age	0.085 ***	0.007	0.073 ***	0.006	0.115 ***	0.015	0.108 ***	0.017
Age Squared	-0.001 ***	0.000	-0.001 ***	0.000	-0.001 ***	0.000	-0.001 ***	0.000
Back	-0.386 ***	0.015	-0.251 ***	0.013	-0.509 ***	0.041	-0.327 ***	0.103
Pardo	-0.301 ***	0.010	-0.200 ***	0.008	-0.383 ***	0.038	-0.277 ***	0.065
Father's Education								
Lower Primary	0.240 ***	0.009	0.069 ***	0.008	0.113 *	0.061	-0.059	0.091
Upper Primary	0.449 ***	0.019	0.110 ***	0.016	0.305	0.204	0.094	0.201
Secondary	0.624 ***	0.022	0.191 ***	0.022	0.588 **	0.292	0.384	0.278
College	0.815 ***	0.028	0.333 ***	0.026	1.255 **	0.515	1.286 ***	0.481
Mother's Education								
Lower Primary	0.276 ***	0.008	0.099 ***	0.008	0.151 **	0.067	-0.012	0.091
Upper Primary	0.519 ***	0.020	0.187 ***	0.019	0.504 ***	0.189	0.286	0.199
Secondary	0.689 ***	0.023	0.286 ***	0.021	0.819 ***	0.279	0.619 **	0.272
College	0.708 ***	0.034	0.288 ***	0.032	0.881 ***	0.306	0.698 **	0.309
Rural	-0.627 ***	0.011	-0.404 ***	0.010	-0.406 ***	0.124	-0.261 **	0.120
Cohort 1933-45	0.005	0.055	0.044	0.044	-0.062	0.071	-0.015	0.072
Cohort 1946-58	0.071 **	0.032	0.063 **	0.025	0.080 **	0.039	0.076 **	0.037
Year 1988	-0.366 ***	0.017	-0.315 ***	0.014	-0.324 ***	0.033	-0.212 ***	0.060
North	-0.215 ***	0.023	-0.150 ***	0.021	-0.177 **	0.083	0.021	0.130
Northeast	-0.300 ***	0.019	-0.213 ***	0.015	-0.229 ***	0.057	-0.033	0.112
South	-0.066 ***	0.022	-0.036 *	0.019	-0.053	0.044	-0.071 *	0.042
Center-West	-0.025	0.020	-0.033 **	0.016	0.135 ***	0.047	0.058	0.061
Constant	3.771 ***	0.147	3.632 ***	0.122	-9.666 ***	2.652	-11.402 ***	2.583
# Obs.	64,375		64,287		63,637		63,549	
R-Squared	0.395		0.505					
Hansen's J-Statistics					14.286		10.464	
Chi-Squared (P-value)					0.354		0.575	

Note: (1) *** statistically significant at 1%;** statistically significant at 5%;* statistically significant at 10%.

(2) The instruments are the nubmer of school by state and year.

Table 8:

Distribution of First Job Occupation if Started to Work 14 Years Old or Below

First Job Occupation	# OBS	Percent
Agriculture	24,978	39.35
Manufacturing	5,910	9.31
Civil Construction	3,109	4.90
Commerce and Transport	6,416	10.11
Services and others	23,059	36.33
Total	63,472	100.00

Table 9:

IV Estimates of Log-Earnings: 25-55 Year-Old Males Pooled 1988 and 1996 - By First Job Occupation Categories

Dependent Variables	Agriculture		Manufacturing		Civil Construction		Commerce and Transport		Services and others	
	Coeff.	Std. Error	Coeff.	Std. Error	Coeff.	Std. Error	Coeff.	Std. Error	Coeff.	Std. Error
Age Started to Work	-2.142	2.319	0.830	0.666	0.775	0.495	0.937 ***	0.298	1.075 **	0.428
Age Started to Work Squared	0.114	0.100	-0.023	0.026	-0.033	0.020	-0.033 ***	0.012	-0.036 **	0.015
Years of Scholing	0.280	0.187	0.147 **	0.066	0.133	0.091	0.147 ***	0.071	-0.083	0.130
Same Father's Occupation	6.775	5.227	-1.061	4.042	-2.260	4.888	-5.188 *	3.009	-4.575	5.775
Interaction Age Sarted to Work and Same Father's Occupation	-0.321	0.471	0.084	0.378	-0.013	0.395	0.448 **	0.209	0.440	0.462
Age	0.055	0.036	0.085 ***	0.020	0.122 **	0.048	0.067 ***	0.029	0.188 ***	0.056
Age Squared	0.000	0.000	-0.001 ***	0.000	-0.001 **	0.001	-0.001	0.000	-0.002 ***	0.001
Back	-0.361	0.329	-0.346 ***	0.072	-0.193 **	0.094	-0.203	0.138	-0.651 ***	0.249
Pardo	-0.132	0.185	-0.131 ***	0.044	-0.211 ***	0.064	-0.095	0.066	-0.542 ***	0.171
Lower Primary	0.000	0.203	-0.118	0.095	0.095	0.125	-0.122	0.127	0.405	0.292
Upper Primary	-0.373	0.450	-0.255	0.157	0.170	0.297	-0.137	0.225	0.697	0.496
Secondary	-1.167	0.907	-0.355	0.218	0.399	0.344	-0.079	0.308	0.971	0.598
College	-2.457 *	1.485	-0.270	0.331	0.502	0.806	0.063	0.401	1.406 **	0.698
Lower Primary	-0.036	0.164	-0.047	0.083	-0.024	0.114	-0.035	0.124	0.408	0.277
Upper Primary	-0.346	0.503	-0.085	0.175	0.200	0.372	0.013	0.225	0.877 *	0.471
Secondary	-0.375	0.709	-0.099	0.196	0.371	0.420	0.043	0.255	1.170 **	0.574
College	-0.344	1.373	-0.054	0.279	-0.006	0.683	0.045	0.284	1.206 **	0.591
Rural	0.199	0.226	-0.035	0.142	-0.287 **	0.129	-0.006	0.183	-0.682 *	0.366
Cohort 1933-45	0.062	0.131	0.008	0.113	0.003	0.212	0.081	0.118	-0.052	0.125
Cohort 1946-58	0.045	0.085	0.059	0.067	0.039	0.106	0.069	0.055	0.033	0.057
Year 1988	-0.362 *	0.205	-0.453 ***	0.105	-0.175	0.185	-0.246 ***	0.055	-0.013	0.183
North	0.211	0.436	-0.118	0.206	0.120	0.181	-0.305 ***	0.071	-0.091	0.127
Northeast	0.577 **	0.264	-0.198	0.165	-0.049	0.116	-0.176 ***	0.067	-0.201 *	0.120
South	1.075 **	0.448	-0.116 **	0.058	-0.066	0.108	-0.013	0.051	-0.115 **	0.048
Center-West	0.238 *	0.128	0.112	0.092	-0.217	0.147	0.042	0.075	0.147	0.099
Constant	10.407	12.877	-3.163	4.342	-1.269	3.250	-2.585	2.179	-4.959	3.147
# Obs.	22,130		6,819		3,451		7,635		23,514	
Hansen's J-Statistics	4.866		9.844		9.756		12.324		2.134	
Chi-Squared (P-value)	0.900		0.454		0.462		0.264		0.995	

Note: (1) *** statistically significant at 1%;** statistically significant at 5%;* statistically significant at 10%.

(2) The instruments are the nubmer of school by state and year.

APPENDIX

Table A1: Summary Statistics for Pooled Sample 1988-1996 - Males Only

Variable	Obs	Mean	Std. Dev.	Min	Max
Log Earnings	95307	5.739453	1.057806	-1.906	10.82
Age Started to Work	94511	12.96044	3.985493	4	48
Years of Schooling	108002	5.886919	4.536432	0	17
Age	108198	37.46757	8.476078	25	55
Black	108186	0.0612556	0.2397997	0	1
Pardo	108186	0.3949679	0.4888461	0	1
Rural	108198	0.1903455	0.3925755	0	1
Father's Education					
Illiterate	73138	0.3716399	0.4832462	0	1
Lower Primary	73138	0.5151084	0.4997751	0	1
Upper Primary	73138	0.047677	0.2130834	0	1
Secondary	73138	0.038106	0.1914536	0	1
College	73138	0.0274686	0.1634456	0	1
Mother's Education					
Illiterate	77225	0.4415798	0.4965786	0	1
Lower Primary	77225	0.4581936	0.4982524	0	1
Upper Primary	77225	0.0483781	0.2145653	0	1
Secondary	77225	0.0409712	0.1982248	0	1
College	77225	0.0108773	0.1037262	0	1
Region 1	108198	0.0245476	0.1547425	0	1
Region 2	108198	0.252574	0.4344906	0	1
Region 3	108198	0.3594983	0.4798556	0	1
Region 4	108198	0.2299858	0.4208253	0	1
Region 5	108198	0.1291336	0.3353493	0	1

Table A2:

First-Stage Regression of the IV estimates From Tables ????: 25-55 Male 1988

	IV 3				IV 4					
Dependent Variables	**Age Started to Work**		**Age Started to Work 2**		**Age Started to Work**		**Age Started to Work 2**		**Schooling**	
	Coeff.	**Std. Error**	**Coeff.**	**Std. Error**	**Coeff.**	**Std. Error**	**Coeff.**	**Std. Error**	**Coeff.**	**Std. Error**
Age	0.016	0.025	1.329	0.734	0.016	0.025	1.328	0.734	0.116	0.024
Age-Squared	0.000	0.000	-0.021	0.010	0.000	0.000	-0.021	0.010	-0.002	0.000
Back	0.073	0.096	-0.023	2.783	0.072	0.096	-0.035	2.783	-1.260	0.092
Pardo	-0.104	0.048	-4.614	1.385	-0.105	0.048	-4.645	1.385	-0.916	0.046
Father's Education										
Lower Primary	0.698	0.053	18.595	1.535	0.698	0.053	18.601	1.535	1.687	0.050
Upper Primary	1.949	0.133	53.710	3.883	1.949	0.133	53.710	3.883	3.855	0.128
Secondary	2.766	0.140	80.632	4.079	2.765	0.140	80.631	4.079	4.998	0.134
College	3.948	0.166	120.874	4.839	3.948	0.166	120.873	4.840	5.632	0.159
Mother's Education										
Lower Primary	0.884	0.051	23.626	1.499	0.884	0.051	23.628	1.499	1.933	0.049
Upper Primary	2.078	0.135	61.926	3.943	2.078	0.135	61.925	3.943	3.902	0.130
Secondary	2.528	0.139	75.432	4.053	2.528	0.139	75.429	4.054	4.520	0.133
College	2.773	0.249	85.620	7.250	2.773	0.249	85.617	7.251	4.483	0.238
Rural	-1.432	0.051	-39.289	1.479	-1.431	0.051	-39.286	1.479	-2.381	0.049
North	0.892	0.095	25.110	2.775	0.893	0.095	25.162	2.776	-0.259	0.091
Northeast	0.412	0.057	11.454	1.673	0.413	0.057	11.488	1.674	-0.555	0.055
South	0.022	0.069	-0.810	2.001	0.023	0.069	-0.799	2.001	-0.053	0.066
Center-West	-0.518	0.090	-13.185	2.611	-0.517	0.090	-13.169	2.612	0.025	0.086
Instruments										
# of School at 6	-0.045	0.055	-1.386	1.600	-0.045	0.055	-1.384	1.600	-0.022	0.053
# of School at 7	0.112	0.120	1.943	3.488	0.111	0.120	1.937	3.488	0.185	0.115
# of School at 8	-0.265	0.170	-4.901	4.944	-0.265	0.170	-4.889	4.944	-0.438	0.163
# of School at 9	0.256	0.172	6.635	5.015	0.256	0.172	6.623	5.015	0.340	0.165
# of School at 10	-0.387	0.176	-11.905	5.137	-0.386	0.176	-11.887	5.138	-0.471	0.169
# of School at 11	0.154	0.171	3.762	4.984	0.154	0.171	3.756	4.984	0.478	0.164
# of School at 12	0.060	0.172	2.390	5.008	0.060	0.172	2.373	5.009	-0.175	0.165
# of School at 13	0.173	0.161	5.252	4.684	0.174	0.161	5.272	4.685	-0.073	0.154
# of School at 14	-0.166	0.149	-4.990	4.340	-0.167	0.149	-5.028	4.341	0.012	0.143
# of School at 15	0.042	0.147	-0.220	4.282	0.044	0.147	-0.185	4.283	0.167	0.141
# of School at 16	0.099	0.137	2.749	4.000	0.099	0.137	2.745	4.001	-0.130	0.132
# of School at 17	-0.029	0.124	1.484	3.611	-0.029	0.124	1.481	3.612	0.243	0.119
# of School at 18	-0.288	0.104	-7.398	3.034	-0.288	0.104	-7.395	3.035	-0.368	0.100
# of School at 19	0.659	0.244	19.842	7.096	0.658	0.244	19.801	7.097	0.610	0.233
# of School at 20	0.013	0.056	0.288	1.620	0.012	0.056	0.272	1.620	-0.041	0.053
Constant	11.276	0.635	112.561	18.499	11.280	0.635	112.679	18.501	2.660	0.608
Obs.	32142		32142		32133		32133		32133	
Test of excluded Instruments										
F(15, OBS-K)	3.480		2.850		3.480		2.850		4.920	
Prob > F	0.000		0.000		0.000		0.000		0.000	
Partial R-squared of Excuded Instruments										
	0.002		0.001		0.002		0.001		0.002	
Shea's Partial R-Squared										

Table A4:

First-Stage Regression of the IV estimates From Tables ????: 25-55 Male 1996

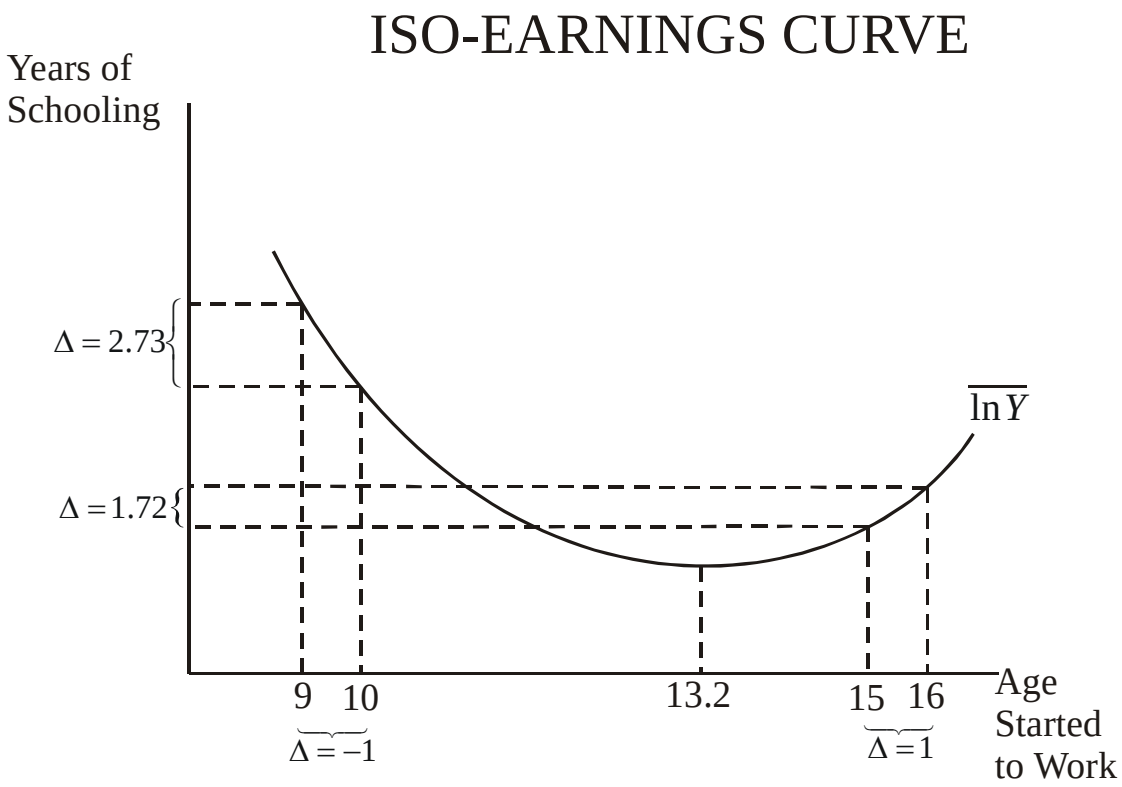
Dependent Variables	IV 3				IV 4					
	Age Started to Work		Age Started to Work 2		Age Started to Work		Age Started to Work 2		Schooling	
	Coeff.	Std. Error	Coeff.	Std. Error	Coeff.	Std. Error	Coeff.	Std. Error	Coeff.	Std. Error
Age	0.060	0.033	1.933	0.955	0.059	0.033	1.912	0.956	0.271	0.032
Age-Squared	-0.001	0.000	-0.031	0.012	-0.001	0.000	-0.031	0.012	-0.004	0.000
Back	0.055	0.088	-0.127	2.546	0.057	0.088	-0.041	2.551	-1.210	0.085
Pardo	-0.169	0.048	-5.915	1.377	-0.169	0.048	-5.907	1.379	-0.999	0.046
Father's Education										
Lower Primary	0.783	0.052	20.745	1.498	0.776	0.052	20.550	1.500	1.823	0.050
Upper Primary	2.253	0.101	63.482	2.932	2.256	0.102	63.606	2.936	3.539	0.098
Secondary	2.818	0.120	80.423	3.478	2.807	0.120	80.159	3.481	4.444	0.117
College	3.984	0.147	122.355	4.236	3.968	0.147	121.943	4.238	5.305	0.142
Mother's Education										
Lower Primary	0.767	0.051	20.354	1.463	0.772	0.051	20.469	1.465	1.691	0.049
Upper Primary	1.582	0.100	44.882	2.891	1.589	0.100	45.069	2.894	3.266	0.097
Secondary	2.302	0.121	68.490	3.484	2.330	0.121	69.240	3.491	4.227	0.117
College	2.255	0.194	66.740	5.608	2.275	0.194	67.338	5.612	4.522	0.188
Rural	-1.593	0.054	-41.769	1.551	-1.595	0.054	-41.814	1.552	-2.376	0.052
North	0.537	0.113	15.540	3.260	0.544	0.113	15.768	3.266	-0.398	0.109
Northeast	0.608	0.061	17.889	1.772	0.606	0.061	17.855	1.773	-0.525	0.059
South	-0.273	0.065	-8.472	1.877	-0.274	0.065	-8.479	1.880	-0.248	0.063
Center-West	-0.593	0.084	-15.000	2.424	-0.596	0.084	-15.073	2.426	-0.062	0.081
Instruments										
# of School at 6	-0.029	0.103	-0.317	2.982	-0.025	0.103	-0.214	2.987	-0.265	0.100
# of School at 7	0.064	0.151	1.109	4.355	0.065	0.151	1.131	4.361	0.152	0.146
# of School at 8	-0.099	0.143	-2.905	4.118	-0.114	0.143	-3.282	4.124	-0.016	0.138
# of School at 9	0.057	0.139	1.491	4.003	0.072	0.139	1.939	4.010	-0.022	0.134
# of School at 10	0.009	0.137	1.000	3.967	0.011	0.137	1.038	3.970	0.124	0.133
# of School at 11	-0.110	0.129	-3.738	3.723	-0.114	0.129	-3.839	3.724	0.051	0.125
# of School at 12	0.019	0.078	0.395	2.253	0.014	0.078	0.257	2.255	0.039	0.076
# of School at 13	-0.003	0.052	0.120	1.515	-0.001	0.052	0.155	1.515	-0.012	0.051
# of School at 14	-0.022	0.045	-0.574	1.292	-0.022	0.045	-0.581	1.292	-0.077	0.043
# of School at 15	-0.063	0.053	-1.156	1.526	-0.064	0.053	-1.173	1.527	-0.045	0.051
# of School at 16	-0.054	0.041	-1.377	1.188	-0.054	0.041	-1.379	1.188	-0.025	0.040
# of School at 17	-0.063	0.046	-1.516	1.322	-0.063	0.046	-1.502	1.322	-0.016	0.044
# of School at 18	-0.078	0.052	-1.797	1.510	-0.077	0.052	-1.774	1.510	-0.116	0.051
# of School at 19	0.020	0.158	1.065	4.559	0.031	0.158	1.354	4.565	0.315	0.153
# of School at 20	-0.012	0.044	-0.094	1.264	-0.012	0.044	-0.074	1.264	-0.033	0.042
Constant	11.706	0.530	136.390	15.330	11.704	0.531	136.338	15.352	0.207	0.514
Obs.	31495		31495		31416		31416		31416	
Test of excluded Instruments										
F(15, OBS-K)	5.170		3.700		5.130		3.680		3.570	
Prob > F	0.000		0.000		0.000		0.000		0.000	
Partial R-squared of Excuded Instruments										
	0.003		0.002		0.002		0.002		0.002	
Shea's Partial R-Squared										
	0.000		0.000		0.000		0.000		0.001	

Table A6:

First-Stage Regression of the IV estimates From Tables ????: 25-55 Male Pooled 1988 and 1996

Dependent Variables	IV 1				IV 2					
	Age Started to Work		Age Started to Work 2		Age Started to Work		Age Started to Work 2		Schooling	
	Coeff.	Std. Error	Coeff.	Std. Error	Coeff.	Std. Error	Coeff.	Std. Error	Coeff.	Std. Error
Age	-0.012	0.022	0.153	0.642	-0.011	0.022	0.169	0.642	0.067	0.021
Age-Squared	0.000	0.000	-0.008	0.008	0.000	0.000	-0.008	0.008	-0.001	0.000
Back	0.071	0.065	0.141	1.881	0.072	0.065	0.185	1.883	-1.229	0.063
Pardo	-0.139	0.034	-5.342	0.976	-0.139	0.034	-5.350	0.977	-0.960	0.032
Father's Education										
Lower Primary	0.743	0.037	19.719	1.072	0.740	0.037	19.627	1.073	1.751	0.036
Upper Primary	2.147	0.081	60.047	2.340	2.150	0.081	60.149	2.343	3.649	0.078
Secondary	2.802	0.091	80.673	2.651	2.797	0.091	80.536	2.653	4.682	0.088
Mother's Education										
College	3.984	0.110	122.231	3.191	3.975	0.110	122.012	3.192	5.461	0.106
Lower Primary	0.822	0.036	21.895	1.047	0.824	0.036	21.955	1.048	1.816	0.035
Upper Primary	1.760	0.080	50.903	2.332	1.764	0.080	51.010	2.333	3.493	0.077
Secondary	2.398	0.091	71.390	2.646	2.413	0.091	71.817	2.649	4.355	0.088
College	2.455	0.153	73.891	4.445	2.466	0.153	74.245	4.447	4.491	0.148
Rural	-1.513	0.037	-40.653	1.068	-1.514	0.037	-40.667	1.069	-2.391	0.035
Cohort 1933-45	0.029	0.102	-0.004	2.950	0.029	0.102	0.030	2.952	-0.245	0.098
Cohort 1946-58	0.068	0.058	1.897	1.681	0.065	0.058	1.849	1.683	0.111	0.056
Year 1988	0.126	0.040	4.357	1.164	0.126	0.040	4.347	1.164	-0.437	0.039
North	0.730	0.072	20.814	2.102	0.735	0.073	20.948	2.104	-0.330	0.070
Northeast	0.475	0.041	13.685	1.200	0.475	0.041	13.691	1.201	-0.577	0.040
South	-0.150	0.047	-5.369	1.352	-0.149	0.047	-5.350	1.353	-0.185	0.045
Center-West	-0.564	0.061	-14.345	1.777	-0.565	0.061	-14.372	1.778	-0.033	0.059
Instruments										
# of School at 6	-0.038	0.047	-1.064	1.366	-0.038	0.047	-1.058	1.366	-0.042	0.045
# of School at 7	0.049	0.088	0.715	2.558	0.052	0.088	0.791	2.559	0.023	0.085
# of School at 8	-0.151	0.108	-2.971	3.134	-0.160	0.108	-3.215	3.136	-0.114	0.104
# of School at 9	0.107	0.105	2.759	3.039	0.115	0.105	2.997	3.042	0.100	0.101
# of School at 10	-0.155	0.105	-4.451	3.037	-0.154	0.105	-4.455	3.040	-0.153	0.101
# of School at 11	0.021	0.094	-0.007	2.738	0.020	0.094	-0.024	2.739	0.144	0.091
# of School at 12	0.097	0.067	2.691	1.955	0.093	0.067	2.590	1.956	0.039	0.065
# of School at 13	0.035	0.048	1.105	1.382	0.037	0.048	1.150	1.382	-0.029	0.046
# of School at 14	-0.004	0.041	-0.290	1.197	-0.004	0.041	-0.293	1.197	-0.053	0.040
# of School at 15	-0.033	0.047	-0.730	1.354	-0.033	0.047	-0.721	1.354	-0.011	0.045
# of School at 16	-0.024	0.038	-0.568	1.111	-0.024	0.038	-0.570	1.111	-0.003	0.037
# of School at 17	-0.055	0.042	-1.044	1.210	-0.055	0.042	-1.033	1.210	0.026	0.040
# of School at 18	-0.128	0.046	-2.849	1.326	-0.128	0.046	-2.843	1.326	-0.152	0.044
# of School at 19	0.283	0.114	8.091	3.296	0.288	0.114	8.214	3.299	0.899	0.110
# of School at 20	-0.018	0.033	-0.290	0.967	-0.018	0.033	-0.294	0.967	-0.012	0.032
Constant	12.428	0.420	153.072	12.202	12.408	0.421	152.599	12.214	3.116	0.406
Obs.	63637		63637		63549		63549		63549	
Test of excluded Instruments										
F(15, OBS-K)	6.810		4.690		6.790		4.680		8.680	
Prob > F	0.000		0.000		0.000		0.000		0.000	
Partial R-squared of Excuded Instruments										
	0.002		0.001		0.002		0.001		0.002	
Shea's Partial R-Squared										
	0.001		0.000		0.000		0.000		0.001	

Figure 1:



Source: Pooled 1988-1996 sample IV2 results for males.