

# Probability Space

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## Stochastic Calculus I

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## Definition

The *sample space*  $\Omega = \{\omega_i : i \in \mathcal{I}\}$  is the set of all possible outcomes of a random experiment. Let  $\mathcal{I}$  be a set of indices. For example,  $\mathcal{I} = \{0, 1, 2, \dots, T\}$ ,  $\mathcal{I} = \{0, 1, 2, \dots\}$ ,  $\mathcal{I} = [0, \infty)$ , etc.

**Example.** If the random experiment consists of rolling a dice, then

$$\Omega = \{ \boxed{1}, \boxed{2}, \boxed{3}, \boxed{4}, \boxed{5}, \boxed{6} \}.$$

## Definition

**(incomplete).** An *event* is a subset of  $\Omega$ .

**Example.**

$$A = \text{the result is even} = \{ \boxed{2}, \boxed{4}, \boxed{6} \}.$$

# Random variables I

## Incomplete definition

### Definition

**(incomplete)**. A *random variable*  $X : \Omega \rightarrow \mathbb{R}$  is a function mapping the sample space (its domain) to real numbers  $\mathbb{R}$ .

Beware! "Sample space" and "random variable" are concepts that should not be confused.

# Random variables II

## Incomplete definition

**Example.** If the random experiment consists of choosing a card at random from a deck of 52 cards, then the event "drawing the king of hearts" is not a random variable, since, among other things, "king of hearts" is not a real number. However, if drawing a certain card is associated with 10 points, then such a relationship is a random variable.

That's why we have chosen to use boxed numbers to denote the possible results of a dice roll: the idea is to distinguish between the event  $\{\boxed{4}\} =$  "the side showing four dots" from the random variable that associates each of the sides of the dice to the number of dots on that side.

**Example.**  $X$ ,  $Y$ ,  $Z$  and  $W$  are random variables:

$\omega$	$X(\omega)$	$Y(\omega)$	$Z(\omega)$	$W(\omega)$
1	0	0	0	5
2	0	5	0	5
3	5	5	0	5
4	5	5	5	5
5	10	5	10	0
6	10	10	10	10

# Probability measures

## Incomplete definition

### Definition

**(incomplete)**.  $\mathbb{P}$  is a *probability measure* on the space  $\Omega$  if :

P1  $\mathbb{P}(\Omega) = 1.$

P2 For any event  $A$  in  $\Omega$ ,  $0 \leq \mathbb{P}(A) \leq 1.$

P3 For any mutually disjoint events  $A_1, A_2, \dots,$   
 $\mathbb{P}(\cup_{i \geq 1} A_i) = \sum_{i \geq 1} \mathbb{P}(A_i)$  where two events  $A_i$  and  $A_j$   
are disjoint if  $A_i \cap A_j = \emptyset.$

## Probability measures

## Example

**Example.** The probability measure  $\mathbb{P}$  represents the situation where the dice is well balanced, while  $\mathbb{Q}$  models a case where the dice is loaded.

$\omega$	$\mathbb{P}(\omega)$	$\mathbb{Q}(\omega)$	$\omega$	$\mathbb{P}(\omega)$	$\mathbb{Q}(\omega)$
1	$\frac{1}{6}$	$\frac{4}{12}$	4	$\frac{1}{6}$	$\frac{1}{12}$
2	$\frac{1}{6}$	$\frac{1}{12}$	5	$\frac{1}{6}$	$\frac{1}{12}$
3	$\frac{1}{6}$	$\frac{1}{12}$	6	$\frac{1}{6}$	$\frac{4}{12}$



## Probability measures

## Equivalent definition

## Theorem

*When  $\text{Card}(\Omega) < \infty$ , say  $\Omega = \{\omega_1, \dots, \omega_n\}$  then the three conditions (P1), (P2) and (P3) in the partial definition of a probability measure are equivalent to the three following conditions:*

$$\mathbf{P1^*} \quad \forall i \in \{1, \dots, n\}, \mathbb{P}(\omega_i) \geq 0.$$

$$\mathbf{P2^*} \quad \text{For any event } A \text{ in } \Omega, \mathbb{P}(A) = \sum_{\omega \in A} \mathbb{P}(\omega).$$

$$\mathbf{P3^*} \quad \sum_{i=1}^n \mathbb{P}(\omega_i) = 1.$$

The proof can be found in the appendix.

## Probability measures

## Properties

## Theorem

*Probability measures have the following properties:*

**P4** For any event  $A$  in  $\Omega$ ,  $\mathbb{P}(A^c) = 1 - \mathbb{P}(A)$ .

**P5**  $\mathbb{P}(\emptyset) = 0$ .

**P6** For any two events  $A$  and  $B$  in  $\Omega$  (not necessarily disjoint),  $\mathbb{P}(A \cup B) = \mathbb{P}(A) + \mathbb{P}(B) - \mathbb{P}(A \cap B)$ .

**P7** If  $A \subseteq B \subseteq \Omega$  then  $\mathbb{P}(A) \leq \mathbb{P}(B)$ .

## Probability measures I

## Proofs

**To be shown:** For any event  $A$  of  $\Omega$ ,  $\mathbb{P}(A^c) = 1 - \mathbb{P}(A)$ .

**Proof of (P4).**

$$1 = \mathbb{P}(\Omega) = \mathbb{P}(A \cup A^c) = \mathbb{P}(A) + \mathbb{P}(A^c)$$

where the first equality comes from (P1) and the third equality comes from (P3). ■

# Probability measures II

## Proofs

**To be shown:**  $\mathbb{P}(\emptyset) = 0$ .

**Proof of (P5).** Property (P5) is nothing more than a special case of (P4): let's replace  $A$  with  $\Omega$ . ■

**To be shown:** *for any two events  $A$  and  $B$  of  $\Omega$  not necessarily disjoint),*

$$\mathbb{P}(A \cup B) = \mathbb{P}(A) + \mathbb{P}(B) - \mathbb{P}(A \cap B).$$

**Proof of (P6).** Since

$$A = (A \cap B) \cup (A \cap B^c) \text{ and } B = (B \cap A) \cup (B \cap A^c)$$

then, using (P3), we get

$$\mathbb{P}(A) = \mathbb{P}(A \cap B) + \mathbb{P}(A \cap B^c) \text{ and } \mathbb{P}(B) = \mathbb{P}(B \cap A) + \mathbb{P}(B \cap A^c).$$

On the other hand,  $A \cup B = (A \cap B^c) \cup (B \cap A^c) \cup (A \cap B)$ , which implies that

$$\begin{aligned} & \mathbb{P}(A \cup B) \\ = & \mathbb{P}(A \cap B^c) + \mathbb{P}(B \cap A^c) + \mathbb{P}(A \cap B) \\ = & [\mathbb{P}(A \cap B^c) + \mathbb{P}(A \cap B)] + [\mathbb{P}(B \cap A^c) + \mathbb{P}(A \cap B)] \\ & - \mathbb{P}(A \cap B) \\ = & \mathbb{P}(A) + \mathbb{P}(B) - \mathbb{P}(A \cap B). \blacksquare \end{aligned}$$

**To be shown:** *If  $A \subseteq B \subseteq \Omega$  then  $\mathbb{P}(A) \leq \mathbb{P}(B)$ .*

**Proof of (P7).** Since  $A \subseteq B$  then  $A \cap B = A$ . Using (P3),

$$\mathbb{P}(B) = \mathbb{P}(A \cap B) + \underbrace{\mathbb{P}(A^c \cap B)}_{\geq 0 \text{ from (P2)}} \geq \mathbb{P}(A \cap B) = \mathbb{P}(A). \blacksquare$$

Probability measures built on  $\Omega$  exist independently from the random variables and vice versa. What is the link between them? That's the topic of the next section.



# Distribution (law)

## Definition

### Definition

The *distribution* or the *law* of a random variable  $X$  is characterized by its (*cumulative*) *distribution function*

$$F_X : \mathbb{R} \rightarrow [0, 1]$$

$x \rightarrow$  probability that the r.v.  $X$  is less than or equal to  $x$ .

So, if  $\mathbb{P}$  is the probability measure assigned to  $\Omega$  then

$$\forall x \in \mathbb{R}, F_X(x) = \mathbb{P} \{ \omega \in \Omega \mid X(\omega) \leq x \}.$$

# Distribution (law)

## Alternative characterization

### Theorem

*In the case where  $\text{Card}(\Omega) < \infty$ , the distribution of a random variable is also characterized by its probability mass function*

$$f_X : \mathbb{R} \rightarrow [0, 1]$$

*$x \rightarrow$  probability that the r.v.  $X$  is equal to  $x$ ,*

*that is*

$$\forall x \in \mathbb{R}, f_X(x) = \mathbb{P}\{\omega \in \Omega \mid X(\omega) = x\}.$$

A proof of this result can be found in the appendix.

## Distribution (law) I

Example

$\omega$	$W(\omega)$	$Q(\omega)$	$\omega$	$W(\omega)$	$Q(\omega)$
1	5	$\frac{4}{12}$	4	5	$\frac{1}{12}$
2	5	$\frac{1}{12}$	5	0	$\frac{1}{12}$
3	5	$\frac{1}{12}$	6	10	$\frac{4}{12}$

## Distribution (law) II

Example

Let's find the probability mass function and the cumulative distribution function of the random variable  $W$ .

$$f_W(x) = \mathbb{Q}\{\omega \in \Omega \mid W(\omega) = x\}$$

$$= \begin{cases} \mathbb{Q}\{\boxed{5}\} & \text{if } x = 0 \\ \mathbb{Q}\{\boxed{1}, \boxed{2}, \boxed{3}, \boxed{4}\} & \text{if } x = 5 \\ \mathbb{Q}\{\boxed{6}\} & \text{if } x = 10 \\ 0 & \text{otherwise} \end{cases}$$

$$= \begin{cases} \frac{1}{12} & \text{if } x = 0 \\ \frac{4}{12} + \frac{1}{12} + \frac{1}{12} + \frac{1}{12} = \frac{7}{12} & \text{if } x = 5 \\ \frac{4}{12} & \text{if } x = 10 \\ 0 & \text{otherwise} \end{cases}$$

## Distribution (law) III

Example

**Recall that:**

$\omega$	$W(\omega)$	$Q(\omega)$	$\omega$	$W(\omega)$	$Q(\omega)$
1	5	$\frac{4}{12}$	4	5	$\frac{1}{12}$
2	5	$\frac{1}{12}$	5	0	$\frac{1}{12}$
3	5	$\frac{1}{12}$	6	10	$\frac{4}{12}$

Let's calculate the cumulative distribution function:

if  $x < 0$  then

$$F_W(x) = \mathbb{Q}\{\omega \in \Omega \mid W(\omega) \leq x\} = \mathbb{Q}(\emptyset) = 0;$$

if  $0 \leq x < 5$  then

$$F_W(x) = \mathbb{Q}\{\omega \in \Omega \mid W(\omega) \leq x\} = \mathbb{Q}\{\boxed{5}\} = \frac{1}{12};$$

if  $5 \leq x < 10$  then

$$F_W(x) = \mathbb{Q}\{\omega \in \Omega \mid W(\omega) \leq x\} = \mathbb{Q}\{\boxed{1}, \boxed{2}, \boxed{3}, \boxed{4}, \boxed{5}\} = \frac{8}{12};$$

if  $x \geq 10$  then

$$F_W(x) = \mathbb{Q}\{\omega \in \Omega \mid W(\omega) \leq x\} = \mathbb{Q}(\Omega) = 1.$$

# Distribution (law)

## Cumulative distribution function properties

- R1 It's a non-decreasing function, that is to say that if  $x < y$  then  $F_X(x) \leq F_X(y)$ .
- R2 It's a function that is right-continuous with left limits.
- R3  $\lim_{x \downarrow -\infty} F_X(x) = 0$  and  $\lim_{x \uparrow \infty} F_X(x) = 1$ .

# Distribution (law)

Notation

**Remark.** In order to alleviate the notation, it is common to write  $\{X \leq x\}$  instead of  $\{\omega \in \Omega \mid X(\omega) \leq x\}$  and  $\mathbb{P}\{X \leq x\}$  instead of  $\mathbb{P}\{\omega \in \Omega \mid X(\omega) \leq x\}$ .



## Distribution (law) I

Example

Sample space

Random  
variablesProbability  
measuresDistributions  
(laws)Sigma-  
algebrasProbability  
measures  
(continued)

References

Appendices

$\omega$	$X(\omega)$	$Y(\omega)$	$Z(\omega)$	$W(\omega)$	$\mathbb{P}(\omega)$	$\mathbb{Q}(\omega)$
1	0	0	0	5	$\frac{1}{6}$	$\frac{4}{12}$
2	0	5	0	5	$\frac{1}{6}$	$\frac{1}{12}$
3	5	5	0	5	$\frac{1}{6}$	$\frac{1}{12}$
4	5	5	5	5	$\frac{1}{6}$	$\frac{1}{12}$
5	10	5	10	0	$\frac{1}{6}$	$\frac{1}{12}$
6	10	10	10	10	$\frac{1}{6}$	$\frac{4}{12}$

## Distribution (law) II

Example

Distributions of the random variables  $X, Y, Z$  and  $W$   
under the probability measure  $\mathbb{P}$

$x$	$\mathbb{P}\{X = x\}$	$\mathbb{P}\{Y = x\}$	$\mathbb{P}\{Z = x\}$	$\mathbb{P}\{W = x\}$
0	$\frac{1}{3}$	$\frac{1}{6}$	$\frac{1}{2}$	$\frac{1}{6}$
5	$\frac{1}{3}$	$\frac{2}{3}$	$\frac{1}{6}$	$\frac{2}{3}$
10	$\frac{1}{3}$	$\frac{1}{6}$	$\frac{1}{3}$	$\frac{1}{6}$

**Remark.** The distribution of the random variable  $X$  is said to be uniform since

$$\mathbb{P}\{X = 0\} = \mathbb{P}\{X = 5\} = \mathbb{P}\{X = 10\} = \frac{1}{3}.$$

## Distribution (law) III

Example

**Recall that:**

Distributions of the random variables  $X, Y, Z$  and  $W$   
under the probability measure  $\mathbb{P}$

$x$	$\mathbb{P}\{X = x\}$	$\mathbb{P}\{Y = x\}$	$\mathbb{P}\{Z = x\}$	$\mathbb{P}\{W = x\}$
0	$\frac{1}{3}$	$\frac{1}{6}$	$\frac{1}{2}$	$\frac{1}{6}$
5	$\frac{1}{3}$	$\frac{2}{3}$	$\frac{1}{6}$	$\frac{2}{3}$
10	$\frac{1}{3}$	$\frac{1}{6}$	$\frac{1}{3}$	$\frac{1}{6}$

**Remark.** The random variables  $Y$  and  $W$  have the same distribution, although they are not equal. Indeed,

$$Y(\boxed{1}) = 0 \neq 5 = W(\boxed{1}).$$

## Distribution (law) IV

Example

Distributions of the random variables  $X, Y, Z$  and  $W$   
under the probability measure  $Q$

$x$	$Q\{X = x\}$	$Q\{Y = x\}$	$Q\{Z = x\}$	$Q\{W = x\}$
0	$\frac{5}{12}$	$\frac{1}{3}$	$\frac{1}{2}$	$\frac{1}{12}$
5	$\frac{2}{12}$	$\frac{1}{3}$	$\frac{1}{12}$	$\frac{7}{12}$
10	$\frac{5}{12}$	$\frac{1}{3}$	$\frac{5}{12}$	$\frac{1}{3}$

**Remark.** Let's note that the distributions of the random variables have changed. Moreover, under the probability measure  $Q$ , the random variables  $Y$  and  $W$  don't have the same distribution any more.

# Random variables

## Equality and equality in law

### Definition

Two random variables  $X$  and  $Y$  are said to be *equal* if and only if  $\forall \omega \in \Omega, X(\omega) = Y(\omega)$ .

They are said to be *equal in distribution* (or in law) when they have the same distribution.

- 1 The concept of equality between two random variables is stronger than the concept of equality in distribution. Indeed, if two random variables are equal, then they are equal in distribution.
- 2 However, two random variables may be equal in distribution but not equal.
- 3 Moreover, two random variables may be equal in distribution under a certain probability measure but not be equal under another probability measure.

In the previous example, when the probability measure  $P$  is assigned to  $\Omega$ ,  $Y$  and  $W$  are equal in distribution but they are not equal.

## Sigma-Algebra

## Introduction

**Question.** If  $\text{Card}(\Omega) = n < \infty$ , how many distinct events are there?

**Answer**  $2^n$ . An event is a subset of  $\Omega = \{\omega_1, \dots, \omega_n\}$ . When we create an event  $A \subseteq \Omega$ , we have, for every  $\omega_i$ , two alternatives: either  $\omega_i \in A$  or  $\omega_i \notin A$ .

**Example.** If  $n = 3$  then

		$\omega_3 \in A$	$\{\omega_1, \omega_2, \omega_3\} = \Omega$
	$\omega_2 \in A$	$\omega_3 \notin A$	$\{\omega_1, \omega_2\}$
$\omega_1 \in A$		$\omega_3 \in A$	$\{\omega_1, \omega_3\}$
	$\omega_2 \notin A$	$\omega_3 \notin A$	$\{\omega_1\}$
	$\omega_2 \in A$	$\omega_3 \in A$	$\{\omega_2, \omega_3\}$
$\omega_1 \notin A$		$\omega_3 \notin A$	$\{\omega_2\}$
	$\omega_2 \notin A$	$\omega_3 \in A$	$\{\omega_3\}$
		$\omega_3 \notin A$	$\emptyset$

# Sigma-Algebra

## Introduction

Usually, we don't need to know the probabilities associated with every event in  $\Omega$ . Such a situation is particularly frequent when  $\text{Card}(\Omega)$  is large or infinite.

## Sigma-Algebra

## Example

**Example.** The sample space is  $\Omega = \{\boxed{1}, \boxed{2}, \boxed{3}, \boxed{4}, \boxed{5}, \boxed{6}\}$ . Let's assume we are interested in the random variable  $X$  only.

$\omega$	$X(\omega)$	$\omega$	$X(\omega)$
$\boxed{1}$	0	$\boxed{4}$	5
$\boxed{2}$	0	$\boxed{5}$	10
$\boxed{3}$	5	$\boxed{6}$	10

The events that characterize the distribution of  $X$  are

$$\{X = 0\} = \{\boxed{1}, \boxed{2}\}; \{X = 5\} = \{\boxed{3}, \boxed{4}\}; \{X = 10\} = \{\boxed{5}, \boxed{6}\}.$$

So, in order to find the distribution of  $X$ , we only need to know the probabilities associated with the events  $\{\boxed{1}, \boxed{2}\}$ ,  $\{\boxed{3}, \boxed{4}\}$  and  $\{\boxed{5}, \boxed{6}\}$ . Knowing, in addition, the probability that event  $\{\boxed{1}\}$  occurs doesn't give us any additional information about the distribution of the random variable  $X$ .



# Sigma-Algebra I

## Definition

The properties of a probability measure are such that, if we know the probability that an event  $A$  occurs, then we also know the probability associated with its complement  $A^c$  since  $\mathbb{P}(A^c) = 1 - \mathbb{P}(A)$ . We are also able to determine the probability associated with the union of a certain number of events characterizing the distribution of  $X$  because of property ( $P6$ ).

## Sigma-Algebra II

## Definition

## Definition

A  $\sigma$ -algebra  $\mathcal{F}$  of  $\Omega$  is a subset of events such that

**T1**  $\Omega \in \mathcal{F}$ .

**T2** If  $A \in \mathcal{F}$  then  $A^c \in \mathcal{F}$ .

**T3** If  $A_1, A_2, \dots \in \mathcal{F}$  then  $\bigcup_{i \geq 1} A_i \in \mathcal{F}$ . In the case where  $\text{Card}(\Omega) < \infty$ , the condition (T3) is equivalent to

(T3\*) If  $A_1, A_2, \dots, A_n \in \mathcal{F}$  then  $\bigcup_{i \geq 1}^n A_i \in \mathcal{F}$ .

Intuitively, the  $\sigma$ -algebra is the set of events in which we are interested.

## Sigma-Algebra

## Examples

- **Example.** The trivial  $\sigma$ -algebra:  $\{\emptyset, \Omega\}$ .
- **Example.** The smallest  $\sigma$ -algebra containing the event  $A$  is  $\{\emptyset, A, A^c, \Omega\}$ .
- **Example.**  $\Omega = \{\boxed{1}, \boxed{2}, \boxed{3}, \boxed{4}, \boxed{5}, \boxed{6}\}$ . The smallest  $\sigma$ -algebra containing the events  $\{\boxed{1}, \boxed{2}\}$ ,  $\{\boxed{3}, \boxed{4}\}$  and  $\{\boxed{5}, \boxed{6}\}$  is

$$\left\{ \emptyset, \{\boxed{1}, \boxed{2}\}, \{\boxed{3}, \boxed{4}\}, \{\boxed{5}, \boxed{6}\}, \{\boxed{1}, \boxed{2}, \boxed{3}, \boxed{4}\}, \right. \\ \left. \{\boxed{1}, \boxed{2}, \boxed{5}, \boxed{6}\}, \{\boxed{3}, \boxed{4}, \boxed{5}, \boxed{6}\}, \Omega \right\}$$

# Sigma-Algebra I

## Definitions

### Definition

The pair  $(\Omega, \mathcal{F})$  made up of a sample space and a  $\sigma$ -algebra is called *measurable space*.

### Definition

A family  $\mathcal{P} = \{A_1, \dots, A_n\}$  of events in  $\Omega$  is called a *finite partition* of  $\Omega$  if

$$(PF1) \quad \forall i \in \{1, \dots, n\}, A_i \neq \emptyset,$$

$$(PF2) \quad \forall i, j \in \{1, \dots, n\} \text{ such that } i \neq j, A_i \cap A_j = \emptyset,$$

$$(PF3) \quad \bigcup_{i=1}^n A_i = \Omega.$$

# Sigma-Algebra II

## Definitions

### Definition

A  $\sigma$ -algebra  $\mathcal{F}$  is said to be generated from the finite partition  $\mathcal{P}$  if it is the smallest  $\sigma$ -algebra that contains all the elements of  $\mathcal{P}$ . In that case  $\mathcal{F}$  is denoted  $\sigma(\mathcal{P})$  and the elements of  $\mathcal{P}$  are the atoms of  $\mathcal{F}$ .

# Random variables

## Definition

### Definition

A random variable  $X$  constructed on the measurable space  $(\Omega, \mathcal{F})$ , is a real-valued function  $X : \Omega \rightarrow \mathbb{R}$  such that

$$(*) \quad \forall x \in \mathbb{R}, \quad \{\omega \in \Omega : X(\omega) \leq x\} \in \mathcal{F}$$

## Random variables

## Definition

**Exercise.** Show that, if  $\text{Card}(\Omega) < \infty$ , then the condition (\*) is equivalent to

$$(**) \forall x \in \mathbb{R}, \{\omega \in \Omega \mid X(\omega) = x\} \in \mathcal{F}.$$

So,  $X : \Omega \rightarrow \mathbb{R}$  is a random variable on the measurable space  $(\Omega, \mathcal{F})$  if and only if the events that characterize its distribution are elements of  $\mathcal{F}$ . If  $\mathcal{F}$  and  $\mathcal{G}$  are two  $\sigma$ -algebras of  $\Omega$  then it is possible that  $X$  is a random variable on the measurable space  $(\Omega, \mathcal{F})$  but that it is not a random variable on the space  $(\Omega, \mathcal{G})$ . To clearly express the fact that the  $\sigma$ -algebra  $\mathcal{F}$  contains the events characterizing the distribution of  $X$ , we say that  $X$  is  $\mathcal{F}$ -measurable.

$$\Omega = \{\boxed{1}, \boxed{2}, \boxed{3}, \boxed{4}, \boxed{5}, \boxed{6}\}$$

$$\text{and } \mathcal{F} = \{\emptyset, \{\boxed{1}, \boxed{3}, \boxed{5}\}, \{\boxed{2}, \boxed{4}, \boxed{6}\}, \Omega\}.$$

The function  $U : \Omega \rightarrow \mathbb{R}$  that returns 1 if the result is even and 0 otherwise is a random variable on  $(\Omega, \mathcal{F})$  whereas the function  $V : \Omega \rightarrow \mathbb{R}$  that returns 1 if the result is less than 4 and 0 otherwise is not. Indeed,

$$\{\omega \in \Omega \mid U(\omega) = x\} = \begin{cases} \{\boxed{1}, \boxed{3}, \boxed{5}\} \in \mathcal{F} & \text{if } x = 0 \\ \{\boxed{2}, \boxed{4}, \boxed{6}\} \in \mathcal{F} & \text{if } x = 1 \\ \emptyset \in \mathcal{F} & \text{otherwise} \end{cases}$$

and

$$\{\omega \in \Omega \mid V(\omega) = x\} = \begin{cases} \{\boxed{4}, \boxed{5}, \boxed{6}\} \notin \mathcal{F} & \text{if } x = 0 \\ \{\boxed{1}, \boxed{2}, \boxed{3}\} \notin \mathcal{F} & \text{if } x = 1 \\ \emptyset \in \mathcal{F} & \text{otherwise} \end{cases}.$$

$U$  is said to be  $\mathcal{F}$ -measurable whereas  $V$  is not.



By contrast,  $U$  and  $V$  are  $\mathcal{G}$ -measurable where

$$\mathcal{G} = \sigma \left\{ \{2\}, \{5\}, \{1, 3\}, \{4, 6\} \right\}$$

$$= \left( \begin{array}{l} \emptyset, \{2\}, \{5\}, \{1, 3\}, \{4, 6\}, \{2, 5\} \\ \{1, 2, 3\}, \{1, 3, 5\}, \{2, 4, 6\}, \{4, 5, 6\}, \\ \{1, 2, 3, 5\}, \{1, 3, 4, 6\}, \{2, 4, 5, 6\} \\ \{1, 3, 4, 5, 6\}, \{1, 2, 3, 4, 6\}, \Omega \end{array} \right)$$

The next results will enable us to identify the smallest  $\sigma$ -algebra that make one or several random variables measurable.

# Random variables

## Measurability

### Theorem

*Let  $(\Omega, \mathcal{F})$ ,  $\text{Card}(\Omega) < \infty$ , be a measurable space and  $\mathcal{P} = \{A_1, \dots, A_n\}$ , be the finite partition of  $\Omega$  that generates  $\mathcal{F}$ . The function  $X : \Omega \rightarrow \mathbb{R}$  is a random variable on that space ( $X$  is  $\mathcal{F}$ -measurable) if and only if  $X$  is constant on the atoms of  $\mathcal{F}$ .*

## Random variables I

Proof

**Proof.** Let's first verify that, *if  $X$  is constant on the atoms of  $\mathcal{F}$  then  $X$  is a random variable on that space  $(\Omega, \mathcal{F})$ .*

If  $X$  is constant on the atoms of  $\mathcal{F}$  then  $X$  may only take a finite number of values that we will denote  $x_1, \dots, x_m$ . So,  $\forall i \in \{1, \dots, m\}$ , the event  $\{\omega \in \Omega \mid X(\omega) = x_i\}$  may be represented as a union of atoms of  $\mathcal{F}$  and, since a  $\sigma$ -algebra is closed under finite unions, then  $\{\omega \in \Omega \mid X(\omega) = x_i\} \in \mathcal{F}$ . ■

## Random variables II

## Proof

Let's now verify that *if  $X$  is a random variable, then  $X$  is constant on the atoms of  $\mathcal{F}$ .*

We're going to use a proof by contradiction. Let's assume that there exists an atom  $A_k$  of  $\mathcal{F}$  on which  $X$  is not constant.

Then there exists at least two values for  $X$  on  $A_k$ . Let  $x_0$ , be one of those values. The event  $A_0 = \{\omega \in \Omega \mid X(\omega) = x_0\}$  is an element of the  $\sigma$ -algebra because, by hypothesis,  $X$  is a random variable. As a result,  $A_0$  is a strict subset of  $A_k$  that belongs to  $\mathcal{F}$ , which contradicts the fact that  $A_k$  is an atom of  $\mathcal{F}$ . ■

The proof of the theorem is complete. ■

# Random variables

## Corollaries

### Corollary

*Any random variable on a measurable space equipped with the trivial  $\sigma$ -algebra is constant.*

### Corollary

*If  $\mathcal{F} = \Omega$  the set of all possible events in  $\Omega$  then any real-valued function on  $\Omega$  ( $X : \Omega \rightarrow \mathbb{R}$ ) is  $\mathcal{F}$ -measurable, that is to say that it is a random variable on  $(\Omega, \mathcal{F})$ .*

# Random variables

## Generated sigma-algebra

### Definition

Let  $X : \Omega \rightarrow \mathbb{R}$ . The smallest  $\sigma$ -algebra  $\mathcal{F}$  that make  $X$  a random variable on the measurable space  $(\Omega, \mathcal{F})$  is called the  $\sigma$ -algebra generated by  $X$  and is denoted  $\sigma(X)$ .

If  $\text{Card}(\Omega) < \infty$  then  $X$  can only take a finite number of values, let's say  $x_1, \dots, x_m$ . For any  $i \in \{1, \dots, m\}$ , let's define  $A_i = \{\omega \in \Omega \mid X(\omega) = x_i\}$ . Then  $\mathcal{P} = \{A_1, \dots, A_m\}$  is a finite partition of  $\Omega$  and  $\sigma(X) = \sigma(\mathcal{P})$ .

# Random variables

## Transformations

### Theorem

*Let's assume that  $\text{Card}(\Omega) < \infty$ . If  $X : \Omega \rightarrow \mathbb{R}$  and  $Y : \Omega \rightarrow \mathbb{R}$  are  $\mathcal{F}$ -measurable, then  $\forall a, b \in \mathbb{R}$ ,  $aX + bY$  is also  $\mathcal{F}$ -measurable, which is to say that any linear combination of random variables built on the same measurable space is a random variable of that space.*

## Random variables

## Transformations

**Proof.** Since  $\text{Card}(\Omega) < \infty$ , the random variables  $X$  and  $Y$  can only take a finite number of values, let's say  $x_1 < \dots < x_m$  and  $y_1 < \dots < y_n$  respectively.  $\forall z \in \mathbb{R}$ ,

$$\begin{aligned} & \{ \omega \in \Omega \mid aX(\omega) + bY(\omega) \leq z \} \\ = & \bigcup_{ax_i + by_j \leq z} \{ \omega \in \Omega \mid X(\omega) = x_i \text{ and } Y(\omega) = y_j \} \\ = & \bigcup_{ax_i + by_j \leq z} \underbrace{\{ \omega \in \Omega \mid X(\omega) = x_i \}}_{\in \mathcal{F}} \cap \underbrace{\{ \omega \in \Omega \mid Y(\omega) = y_j \}}_{\in \mathcal{F}} \in \mathcal{F}. \blacksquare \end{aligned}$$



## Probability measures

## Definitions

## Definition

Let  $(\Omega, \mathcal{F})$  be a measurable space. The function  $\mathbb{P} : \mathcal{F} \rightarrow [0, 1]$  is a *probability measure* on  $(\Omega, \mathcal{F})$  if

**P1**  $\mathbb{P}(\Omega) = 1.$

**P2**  $\forall A \in \mathcal{F}, 0 \leq \mathbb{P}(A) \leq 1.$

**P3**  $\forall A_1, A_2, \dots \in \mathcal{F}$  such that  $A_i \cap A_j = \emptyset$  if  $i \neq j$ ,  
 $\mathbb{P}(\bigcup_{i \geq 1} A_i) = \sum_{i \geq 1} \mathbb{P}(A_i).$

## Definition

The triple  $(\Omega, \mathcal{F}, \mathbb{P})$  made up of a sample space, a  $\sigma$ -algebra on  $\Omega$  and a probability measure built on  $(\Omega, \mathcal{F})$  is called *probability space*.

## References

- 1 BILLINGSLEY, Patrick (1986). *Probability and Measure*, Second Edition, Wiley, New York.
  - That book is meant for people with a strong background in mathematics. All important results in probability theory can be found in that book with their proofs. The book offers a wide selection of exercises, at all difficulty levels. That book is well above the level set for this course.
- 2 ROSS, Sheldon, M. (1997). *Introduction to probability Models*, sixth edition, Academic Press, New York.

# Probability measures I

## Appendix I

### Theorem

*When  $\text{Card}(\Omega) < \infty$ , let's say  $\Omega = \{\omega_1, \dots, \omega_n\}$  then the three conditions (P1), (P2) and (P3) of the partial definition of a probability measure are equivalent to the following three conditions:*

### Definition

$$\text{P1}^* \quad \forall i \in \{1, \dots, n\}, \mathbb{P}(\omega_i) \geq 0.$$

$$\text{P2}^* \quad \text{For any event } A \text{ in } \Omega, \mathbb{P}(A) = \sum_{\omega \in A} \mathbb{P}(\omega).$$

$$\text{P3}^* \quad \sum_{i=1}^n \mathbb{P}(\omega_i) = 1.$$

**Proof.** Let's first assume that  $\mathbb{P}$  satisfies the three conditions  $(P1)$ ,  $(P2)$  and  $(P3)$  of the definition of a probability measure and let's show that, then,  $\mathbb{P}$  also satisfies the conditions  $(P1^*)$ ,  $(P2^*)$  and  $(P3^*)$ .

- $(P1^*)$  Since  $\forall i \in \{1, \dots, n\}$ ,  $\{\omega_i\}$  is an event in  $\Omega$  then, from condition  $(P2)$ ,  $0 \leq \mathbb{P}(\omega_i) \leq 1$ .
- $(P2^*)$  Since  $\{\omega_1\}, \dots, \{\omega_n\}$  are mutually disjoint events, then, using condition  $(P3)$ , we obtain that, for any event  $A$  in  $\Omega$ ,  $\mathbb{P}(A) = \sum_{\omega \in A} \mathbb{P}(\omega)$ .
- $(P3^*)$  Since  $\Omega$  is itself an event, by using the equality that we have just established and property  $(P1)$ , we have that  $\sum_{i=1}^n \mathbb{P}(\omega) = \mathbb{P}(\Omega) = 1$ .

Let's now assume that  $\mathbb{P}$  satisfies conditions  $(P1^*)$ ,  $(P2^*)$  and  $(P3^*)$  and let's show that, then,  $\mathbb{P}$  also satisfies the three conditions of the partial definition of a probability measure.

- $(P1)$  Using successively  $(P2^*)$  and  $(P3^*)$ ,  

$$\mathbb{P}(\Omega) = \sum_{i=1}^n \mathbb{P}(\omega_i) = 1.$$
- $(P2)$  Since condition  $(P1^*)$  implies that  $\forall \omega \in \Omega$ ,  
 $\mathbb{P}(\omega) \geq 0$ , then for any event  $A$  in  $\Omega$ ,  

$$\mathbb{P}(A) = \sum_{\omega \in A} \mathbb{P}(\omega) \geq 0$$
 where the equality is warranted by  $(P2^*)$ .

On the other hand, by using successively  $(P2^*)$ , the inequality above and  $(P3^*)$ , 
$$\mathbb{P}(A) = \sum_{\omega \in A} \mathbb{P}(\omega) \leq \sum_{\omega \in A} \mathbb{P}(\omega) + \underbrace{\sum_{\omega \in A^c} \mathbb{P}(\omega)}_{\geq 0} = \sum_{\omega \in \Omega} \mathbb{P}(\omega) = 1.$$

## Probability measures IV

## Appendix I

- (P3) For all events  $A_1, A_2, \dots$  mutually disjoint,

$$\begin{aligned}\mathbb{P}\left(\bigcup_{i \geq 1} A_i\right) &= \mathbb{P}\left(\bigcup_{\omega \in \bigcup_{i \geq 1} A_i} \omega\right) \\ &= \sum_{\omega \in \bigcup_{i \geq 1} A_i} \mathbb{P}(\omega) \\ &= \sum_{i \geq 1} \sum_{\omega \in A_i} \mathbb{P}(\omega) \\ &= \sum_{i \geq 1} \mathbb{P}(A_i). \blacksquare\end{aligned}$$

## Probability measures I

## Appendix II

## Theorem

*In the case where  $\text{Card}(\Omega) < \infty$ , the distribution of a random variable is also characterized by its probability mass function*

$$f_X : \mathbb{R} \rightarrow [0, 1]$$

*$x \rightarrow$  probability that the r.v.  $X$  is equal to  $x$ ,*

*which is to say that*

$$\forall x \in \mathbb{R}, f_X(x) = \mathbb{P}\{\omega \in \Omega \mid X(\omega) = x\}.$$

**Proof.** We must prove that

- (i) Given a probability mass function  $f_X$ , there exists one and only one cumulative distribution function,
- (ii) for any given cumulative distribution function  $F_X$ , there exists one and only one probability mass function.

Since  $\text{Card}(\Omega) < \infty$  then the random variable  $X$  can only take a finite number of values, let's say  $x_1 < \dots < x_n$ .



**Proof of point (i).** Let  $f_X$  be the probability mass function of the random variable  $X$ . Then,  $\forall x \in \mathbb{R}$ ,

$$\begin{aligned} F_X(x) &= \mathbb{P}\{\omega \in \Omega \mid X(\omega) \leq x\} \\ &= \mathbb{P}\left[\bigcup_{x_i \leq x} \{\omega \in \Omega \mid X(\omega) = x_i\}\right] \\ &= \sum_{x_i \leq x} \mathbb{P}\{\omega \in \Omega \mid X(\omega) = x_i\} \\ &\quad \text{because the events in question are disjoint.} \\ &= \sum_{x_i \leq x} f_X(x_i). \end{aligned}$$

Since, by construction, the probability mass function is unique, then, the cumulative distribution function is also unique.

**Proof of point (ii).** Let  $F_X$ , be the cumulative distribution function of the random variable  $X$ . Then,  $\forall i \in \{2, \dots, n\}$ ,

$$\begin{aligned} & F_X(x_i) \\ &= \mathbb{P} \{ \omega \in \Omega \mid X(\omega) \leq x_i \} \\ &= \mathbb{P} \{ \{ \omega \in \Omega \mid X(\omega) = x_i \} \cup \{ \omega \in \Omega \mid X(\omega) \leq x_{i-1} \} \} \\ &= \mathbb{P} \{ \omega \in \Omega \mid X(\omega) = x_i \} + \mathbb{P} \{ \omega \in \Omega \mid X(\omega) \leq x_{i-1} \} \\ & \quad \text{because both events in question are disjoint.} \\ &= f_X(x_i) + F_X(x_{i-1}). \end{aligned}$$

Which implies that  $\forall i \in \{2, \dots, n\}$ ,

$$f_X(x_i) = F_X(x_i) - F_X(x_{i-1}).$$

What happens to  $x_1$ ? Since  $x_1$  is the smallest possible value for the random variable  $X$ ,

$$\begin{aligned}f_X(x_1) &= \mathbb{P}\{\omega \in \Omega \mid X(\omega) = x_1\} \\ &= \mathbb{P}\{\omega \in \Omega \mid X(\omega) \leq x_1\} \\ &= F_X(x_1).\end{aligned}$$

Now,  $\forall x \in \mathbb{R}, x \notin \{x_1, \dots, x_n\}$ ,

$$f_X(x) = \mathbb{P}\{\omega \in \Omega \mid X(\omega) = x\} = \mathbb{P}\{\emptyset\} = 0.$$

Since, by construction, the cumulative distribution function is unique, then, the probability mass function is also unique. ■  
That last argument completes the proof of the proposition. ■